

## The Normal Distribution Curve is a P.D.F.

### Gamma Function/Distribution

Let  $f(y) = \begin{cases} ky^{\alpha-1}e^{-y/\beta} & \text{for } y > 0 \\ 0 & \text{for } y \leq 0 \end{cases}$  where  $\alpha > 0, \beta > 0$ . For  $f$  to be a valid p.d.f. it

must be that  $\int_0^{\infty} f(y)dy = 1$ .

$\Rightarrow k \int_0^{\infty} y^{\alpha-1} e^{-y/\beta} dy = 1$ . Now, let  $x = \frac{y}{\beta}$ ,  $\beta dx = dy$  so that the equation becomes

$$k \int_0^{\infty} (\beta x)^{\alpha-1} e^{-x} \beta dx = 1$$

$$k \beta^{\alpha} \int_0^{\infty} x^{\alpha-1} e^{-x} dx = 1$$

$k \beta^{\alpha} \Gamma(\alpha) = 1$  where  $\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx =$  the *Gamma Function* for  $\alpha > 0$ .

Now, if we set  $k = \frac{1}{\beta^{\alpha} \Gamma(\alpha)}$  we get the p.d.f. of the *Gamma Distribution*:

$$f(y) = \begin{cases} \frac{1}{\beta^{\alpha} \Gamma(\alpha)} y^{\alpha-1} e^{-y/\beta} & \text{for } y > 0 \\ 0 & \text{for } y \leq 0 \end{cases}$$

Fact:  $\Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1)$  for  $\alpha > 1$ .

$$\text{Proof: } \Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx = \lim_{m \rightarrow \infty} \int_0^m x^{\alpha-1} e^{-x} dx$$

Let  $I = \int x^{\alpha-1} e^{-x} dx$  and use *Integration by Parts* with:  $u = x^{\alpha-1}$   $dv = e^{-x} dx$  to  
 $du = (\alpha - 1)x^{\alpha-2} dx$   $v = -e^{-x}$

obtain  $I = -x^{\alpha-1} e^{-x} + (\alpha - 1) \int x^{\alpha-2} e^{-x} dx$  and so

$$\Gamma(\mathbf{a}) = \lim_{m \rightarrow \infty} \left( -x^{a-1} e^{-x} \Big|_0^m \right) + (\mathbf{a} - 1) \int_0^{\infty} x^{a-2} e^{-x} dx \spadesuit$$

$$= 0 + (\alpha - 1)\Gamma(\alpha - 1) = (\alpha - 1)\Gamma(\alpha - 1) \quad \text{Q.E.D.}$$

Now, if  $\mathbf{a} = 1$  we get  $\Gamma(1) = \int_0^{\infty} e^{-y} dy = 1$  (Check!!)

so that  $\tilde{\Lambda}(1) = 1$ . It follows with the result before that if  $n$  is a Natural number then:

$$\begin{aligned} \tilde{\Lambda}(n) &= (n - 1)\tilde{\Lambda}(n - 1) \\ &= (n - 1)(n - 2)\tilde{\Lambda}(n - 2) \\ &= (n - 1)(n - 2)(n - 3)\tilde{\Lambda}(n - 3) \end{aligned}$$

and eventually you arrive at:

$$\tilde{\Lambda}(n) = (n - 1)(n - 2)(n - 3) \dots (3)(2)\tilde{\Lambda}(1) = (n - 1)!$$

That is:  $\tilde{\Lambda}(n) = (n - 1)!$ , or, if you prefer,  $n! = \tilde{\Lambda}(n + 1)$ .

*If  $n \in \mathbb{N}$  then  $\Gamma(n) = (n - 1)!$*

Consider the integral  $2^{1-a} \int_0^{\infty} z^{2a-1} e^{-\frac{1}{2}z^2} dz$  for  $\mathbf{a} > 0$  and use the substitution

$z^2 = 2x$ ,  $zdz = dx$  to obtain:

$$\begin{aligned} 2^{1-a} \int_0^{\infty} z^{2a-1} e^{-\frac{1}{2}z^2} dz &= 2^{1-a} \int_0^{\infty} z^{2a-2} e^{-\frac{1}{2}z^2} z dz = 2^{1-a} \int_0^{\infty} (z^2)^{a-1} e^{-\frac{1}{2}z^2} z dz \\ &= 2^{1-a} \int_0^{\infty} (2x)^{a-1} e^{-\frac{1}{2}2x} dx = 2^{1-a} 2^{a-1} \int_0^{\infty} x^{a-1} e^{-x} dx \\ &= \int_0^{\infty} x^{a-1} e^{-x} dx = \Gamma(\alpha) \end{aligned}$$

So,  $2^{1-a} \int_0^{\infty} z^{2a-1} e^{-\frac{1}{2}z^2} dz = \Gamma(\mathbf{a})$

Immediate consequence of the above is that

$$\Gamma(1/2) = 2^{1-\frac{1}{2}} \int_0^{\infty} z^{2\frac{1}{2}-1} e^{-\frac{1}{2}z^2} dz = \sqrt{2} \int_0^{\infty} e^{-\frac{1}{2}z^2} dz \quad (*)$$

Therefore,  $(\Gamma(1/2))^2 = 2 \left( \int_0^{\infty} e^{-\frac{1}{2}z^2} dz \right)^2 = 2 \int_0^{\infty} e^{-\frac{1}{2}x^2} dx \int_0^{\infty} e^{-\frac{1}{2}y^2} dy$

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$\spadesuit \lim_{m \rightarrow \infty} \left( -x^{a-1} e^{-x} \Big|_0^m \right) = 0$  after “sufficient” applications of L’Hospital’s Rule.

$$= 2 \int_0^{\infty} \int_0^{\infty} e^{-\frac{1}{2}x^2} e^{-\frac{1}{2}y^2} dx dy = 2 \int_0^{\infty} \int_0^{\infty} e^{-\frac{1}{2}(x^2+y^2)} dx dy$$

Change to polar coordinates by using  $x = r \cos \theta$ ,  $y = r \sin \theta$ ,  $dx dy = r dr d\theta$  to obtain

$$= 2 \int_{\theta=0}^{\pi/2} \int_{r=0}^{\infty} r e^{-\frac{1}{2}r^2} dr d\theta = 2 \int_{\theta=0}^{\pi/2} d\theta = \pi$$

Therefore,  $(\Gamma(1/2))^2 = \pi$  so  $\Gamma(1/2) = \sqrt{\pi}$  !!!

### **The P. D. F. of the Normal Distribution**

Consider the function  $n(y) = \frac{1}{s\sqrt{2p}} e^{-\frac{1}{2}\left(\frac{y-m}{s}\right)^2}$ .

$$\int_{-\infty}^{\infty} n(y) dy = \int_{-\infty}^{\infty} \frac{1}{s\sqrt{2p}} e^{-\frac{1}{2}\left(\frac{y-m}{s}\right)^2} dy \quad (\text{Let } z = \frac{y-m}{s}, dz = \frac{1}{s} dy)$$

$$= \frac{1}{\sqrt{2p}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}z^2} dz = \frac{2}{\sqrt{2p}} \int_0^{\infty} e^{-\frac{1}{2}z^2} dz$$

$$= \frac{1}{\sqrt{p}} \cdot \sqrt{2} \int_0^{\infty} e^{-\frac{1}{2}z^2} dz = \frac{1}{\sqrt{p}} \Gamma(1/2) \quad (\text{by } (*))$$

$$= \frac{1}{\sqrt{p}} \cdot \sqrt{p} = 1$$

Therefore,  $n(y) = \frac{1}{s\sqrt{2p}} e^{-\frac{1}{2}\left(\frac{y-m}{s}\right)^2}$  is a valid *probability density function* – the p. d. f. of the *Normal Distribution*.